

**ECEn 670**

**Homework Problem Set 6**

Due at beginning of class, Thursday, November 19, 2009

Problems are from *An Introduction to Statistical Signal Processing* by Gray and Davisson unless otherwise specified.

1. 5.1
2. 5.7
3. 5.13
4. 5.14
5. 5.15
6. 5.29
7. 5.30

5.1)  $\bar{X}_n$ : iid Gaussian with mean  $m$  and variance  $\sigma^2$ .

$h$ :  $\delta$ -response,  $h_0=1$ ,  $h_1=r$ ,  $h_k=0$  for other  $k$ .

$$W_n = \bar{X}_n + r\bar{X}_{n-1}$$

Two-sided:

$$\begin{aligned} E W_n &= E[\bar{X}_n + r\bar{X}_{n-1}] = E\bar{X}_n + r E\bar{X}_{n-1} \\ &= (1+r)m \end{aligned}$$

$$\begin{aligned} R_W(k,j) &= E[(\bar{X}_k + r\bar{X}_{k-1})(\bar{X}_j + r\bar{X}_{j-1})] \\ &= E[\bar{X}_k \bar{X}_j] + r E[\bar{X}_k \bar{X}_{j-1}] + r E[\bar{X}_{k-1} \bar{X}_j] + r^2 E[\bar{X}_{k-1} \bar{X}_{j-1}] \\ &= \begin{cases} (1+r)^2 m^2 + (1+r^2) \sigma^2 & |k-j|=0 \\ (1+r)^2 m^2 + r^2 \sigma^2 & |k-j|=1 \\ (1+r)^2 m^2 & |k-j| > 1 \end{cases} \end{aligned}$$

This process,  $\{W_n\}$ , is wide-sense stationary. Because  $W_n$  is the sum of Gaussians, it too is Gaussian. Because  $\{W_n\}$  is wide-sense stationary and Gaussian, it is strictly stationary.

One-sided:

$$E(W_n) = \begin{cases} E(\bar{X}_0) = m, & n=0 \\ E(\bar{X}_n + r\bar{X}_{n-1}) = (1+r)m, & n > 1 \end{cases}$$

$$\begin{aligned} R_W(k,j) &= \begin{cases} E[\bar{X}_0^2] = \sigma^2 + m^2, & k=0, j=0 \\ E[\bar{X}_0(\bar{X}_1 + r\bar{X}_0)] = m^2 + r(\sigma^2 + m^2), & k=0, j=1 \text{ or } k=1, j=0 \\ E[\bar{X}_0(\bar{X}_n + r\bar{X}_{n-1})] = (1+r)m^2, & k=0, j>1, \text{ or } k>1, j=0 \\ \text{Same as two sided} & k>1, j>1 \end{cases} \end{aligned}$$

$$\lim_{n \rightarrow \infty} E W_n = (1+r)m$$

$$\lim_{n \rightarrow \infty} R_W(n, n+k) = \begin{cases} (1+r)^2 m^2 + (1+r^2) \sigma^2, & k=0 \\ (1+r)^2 m^2 + r^2 \sigma^2, & |k|=1 \\ (1+r)^2 m^2, & |k| > 1 \end{cases}$$

5.7)  $\{X_n\}$  is iid Gaussian with zero mean and  $R_X(0) = \sigma^2$ .

$\{U_n\}$  iid binary, independent of  $X$ , with  $\Pr(U_n=1) = \Pr(U_n=-1) = 1/2$ .

$$Z_n = X_n U_n$$

$$Y_n = U_n + X_n$$

$$W_n = U_0 + X_n, \text{ all } n.$$

Mean  $E[Z_n] = E[X_n] E[U_n] = 0$

Covariance  $K_Z(k, j) = \begin{cases} E[(Z_k - E(Z_k))^2] = E[X_n^2] E[U_n^2] = \sigma^2(1) = \sigma^2, & k=j \\ 0 & , k \neq j \end{cases}$

PSD  $S_Z(f) = \sigma^2$  for all  $f$

Mean  $E[Y_n] = E[U_n] + E[X_n] = 0$ .

Covariance  $K_Y(k, j) = \begin{cases} E[U_k^2] + E[X_k^2] + 2E[U_k X_k] = \sigma^2 + 1 + 0 = \sigma^2 + 1, & k=j \\ 0 & , k \neq j \end{cases}$

PSD  $S_Y(f) = 1 + \sigma^2$  for all  $f$ .

Mean  $E[W_n] = E[U_0 + X_n] = E[U_0] + E[X_n] = 0$

Covariance  $K_W(k, j) = E[(U_0 + X_k)(U_0 + X_j)] = E[U_0^2] + E[X_k X_j] + 2E[U_0] E[X_k]$   
 $= 1 + \sigma^2 \delta_{k,j}$   
 $K_W(k) = 1 + \sigma^2 \delta_k$

PSD  $S_W(f) = \delta(f) + \sigma^2$

Cross covariances

$$K_{ZY} = E[Z_n Y_n] = E[X_n U_n (U_n + X_n)] = E[X_n U_n^2 + U_n X_n^2] = 0$$

$$K_{ZW} = E[Z_n W_n] = E[X_n U_n (U_0 + X_n)] = E[X_n U_n U_0 + X_n^2 U_0] = 0$$

$$K_{YW} = E[(U_n + X_n)(U_0 + X_m)] = E[U_n U_0 + X_n U_0 + U_n X_m + X_n X_m]$$
  
 $= 1 \delta_n + \sigma^2 \delta_{n-m}$

$$5.13) S_{\underline{X}}(f) = \frac{N_0}{2} \xrightarrow{\mathcal{F}} R_{\underline{X}}(t) = \frac{N_0}{2} \delta(t)$$

To have a PSD, it must be weakly stationary.

$$h(t) \xleftrightarrow{\mathcal{F}} H(f)$$

$$g(t) \xleftrightarrow{\mathcal{F}} G(f)$$

$$\underline{Y}(t) = \int_0^{\infty} h(\tau) \underline{X}(t-\tau) d\tau$$

$$\underline{V}(t) = \int_0^{\infty} g(\tau) \underline{X}(t-\tau) d\tau$$

$$a) R_{\underline{Y}, \underline{V}}(t, s) = E(\underline{Y}_t V_s)$$

$$= E\left[ \int_0^{\infty} h(\tau_t) \underline{X}(t-\tau_t) d\tau_t \int_0^{\infty} g(\tau_s) \underline{X}(s-\tau_s) d\tau_s \right]$$

$$= E\left[ \int_0^{\infty} \int_0^{\infty} h(\tau_t) g(\tau_s) \underline{X}(t-\tau_t) \underline{X}(s-\tau_s) d\tau_t d\tau_s \right]$$

$$= \int_0^{\infty} \int_0^{\infty} h(\tau_t) g(\tau_s) E[\underline{X}(t-\tau_t) \underline{X}(s-\tau_s)] d\tau_t d\tau_s$$

$$= \int_0^{\infty} \int_0^{\infty} h(\tau_t) g(\tau_s) R_{\underline{X}}(t-\tau_t, s-\tau_s) d\tau_t d\tau_s$$

$$= \int_0^{\infty} \int_0^{\infty} h(\tau_t) g(\tau_s) R_{\underline{X}}(t-\tau_t-s+\tau_s) d\tau_t d\tau_s$$

$$= \frac{N_0}{2} \int_0^{\infty} \int_0^{\infty} h(\tau_t) g(\tau_s) \delta(t-\tau_t-s+\tau_s) d\tau_t d\tau_s$$

$$= \frac{N_0}{2} \int_0^{\infty} h(\tau_t) g(s-t+\tau_t) d\tau_t$$

$$= \frac{N_0}{2} \int_0^{\infty} h(\tau_t) g(\tau_t - (t-s)) d\tau_t$$

If  $E(\underline{Y}_t V_s) = 0$ , then  $\underline{Y}_t$  and  $V_s$  are orthogonal, so uncorrelated, and because they are Gaussian, then independent. This integral is the convolution of  $h$  and time-reversed  $g$ , which results in a Fourier transform of  $H(f)G(-f)$ .

$$\int d\tau_t h(\tau_t) g(\tau_t - (t-s)) = \int H(f) G(-f) e^{j2\pi f(t-s)} df$$

This will be zero if  $H(f)G(-f) = 0$  for all  $f$ .

You can see this is the case if you consider that if any spectral components are shared between  $H(f)$  and  $G(-f)$ , then there will be information that is shared between  $\underline{Y}(t)$  and  $\underline{V}(t)$ .

5.14  $\{X(t)\}$  and  $\{Y(t)\}$  be zero-mean stationary Gaussian with  $R(\tau)$  and  $S(f)$ .

$X(t)$  and  $Y(t)$  independent

$$E[X(t)Y(s)] = 0, \text{ all } t, s.$$

$$\sigma^2 = R(0).$$

$$W(t) = X(t) \cos(2\pi f_0 t) + Y(t) \sin(2\pi f_0 t)$$

$$E[W(t)] = E[X(t) \cos(2\pi f_0 t) + Y(t) \sin(2\pi f_0 t)]$$

$$= E[X(t)] \cos(2\pi f_0 t) + E[Y(t)] \sin(2\pi f_0 t)$$

$$= 0$$

$$R_W(t, s) = E[W(t)W(s)]$$

$$= E[(X(t) \cos(2\pi f_0 t) + Y(t) \sin(2\pi f_0 t))(X(s) \cos(2\pi f_0 s) + Y(s) \sin(2\pi f_0 s))]$$

$$= E[(X(t)X(s))] \cos(2\pi f_0 t) \cos(2\pi f_0 s) + 0 + 0 + E[Y(t)Y(s)] \sin(2\pi f_0 t) \sin(2\pi f_0 s)$$

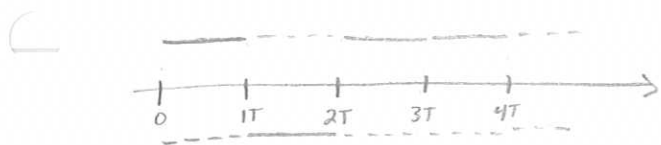
$$= R(t-s) [\cos(2\pi f_0 t) \cos(2\pi f_0 s) + \sin(2\pi f_0 t) \sin(2\pi f_0 s)]$$

$$= R(t-s) [\cos(2\pi f_0 (t-s))]$$

$\{W(t)\}$  is weakly stationary because the mean and correlation only depend on the time difference.

### 5.15) PAM Process

$$X(t) = X_n; t \in [(n-1)T, nT]$$



Example waveform for PAM

$$X(t) = \sum_k X_k p(t-kT)$$

$$Y(t) = X(t+Z) \text{ where } Z \text{ is random variable, uniformly distributed on } [0, T]$$

$$E[Y(t)] = E[X(t+Z)] = 1(\frac{1}{2}) + (-1)(\frac{1}{2}) = 0$$

$$R_X(t, s) = E[Y(t)Y(s)] = E[X(t+Z)X(s+Z)]$$

$$\text{If } |t-s| > T, \text{ then } E[X(t+Z)X(s+Z)] = \frac{1}{4}(1)(1) + \frac{1}{4}(1)(-1) + \frac{1}{4}(-1)(1) + \frac{1}{4}(-1)(-1) = 0$$

If  $|t-s| \leq T$ , we need to know the probability of overlap. When they overlap  $X(t+Z)X(s+Z) = 1$ . They are both either one or negative one.

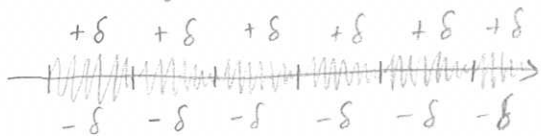
$$Pr(\text{overlap}) = 1 - \frac{|t-s|}{T}$$

$$\therefore E[X(t+Z)X(s+Z)] = (1)(1 - \frac{|t-s|}{T}) = 1 - \frac{|t-s|}{T}$$

$$R_X(t, s) = \begin{cases} 1 - \frac{|t-s|}{T}, & |t-s| \leq T \\ 0, & |t-s| > T \end{cases}$$

### PSK Process

$$U(t) = a_0 \cos(2\pi f_0 t + \delta X(t))$$



The phase is shifted in each time block

$$V(t) = U(t+\Theta) \text{ where } \Theta \text{ is a random variable uniformly distributed on } [0, \frac{1}{f_0}]$$

$$E[V(t)] = E[U(t+\Theta)] = E[a_0 \cos(2\pi f_0(t+\Theta) + \delta X(t+\Theta))]$$

$$= a_0 E[\cos(2\pi f_0(t+\Theta)) \cos(\delta X(t+\Theta)) - \sin(2\pi f_0(t+\Theta)) \sin(\delta X(t+\Theta))]$$

If  $X(t+\Theta)$  stays the same throughout, then

$$\int_0^{1/f_0} \cos(2\pi f_0(t+\Theta)) d\Theta = 0$$

If  $\mathcal{X}(t+\theta)$  changes, then it is equally likely to go from +1 to -1 as it is to go from -1 to +1.

Thus, even function so stays constant

$$\begin{aligned} & a_0 E \left[ \cos(2\pi f_0(t+\theta)) \overbrace{\cos(\delta \mathcal{X}(t+\theta))}^{\text{even function so stays constant}} - \sin(2\pi f_0(t+\theta)) \sin(\delta \mathcal{X}(t+\theta)) \right] \\ &= \frac{1}{2} \left[ \int_0^a -\sin(2\pi f_0(t+\theta)) \sin(\delta) d\theta + \int_a^{1/f_0} -\sin(2\pi f_0(t+\theta)) \sin(-\delta) d\theta \right] \\ &+ \frac{1}{2} \left[ \int_0^a -\sin(2\pi f_0(t+\theta)) \sin(-\delta) d\theta + \int_a^{1/f_0} -\sin(2\pi f_0(t+\theta)) \sin(\delta) d\theta \right] \\ &= 0 \end{aligned}$$

$$\therefore E[V(t)] = 0.$$

$$\begin{aligned} R_V(t,s) &= E[V(t)V(s)] = E[U(t+\theta)U(s+\theta)] \\ &= E \left[ a_0 \cos(2\pi f_0(t+\theta) + \delta \mathcal{X}(t+\theta)) a_0 \cos(2\pi f_0(s+\theta) + \delta \mathcal{X}(s+\theta)) \right] \\ &= \frac{a_0^2}{2} E \left[ \cos(2\pi f_0(t+\theta) - 2\pi f_0(s+\theta) + \delta \mathcal{X}(t+\theta) - \delta \mathcal{X}(s+\theta)) \right. \\ &\quad \left. + \cos(2\pi f_0(t+\theta + s+\theta) + \delta \mathcal{X}(t+\theta) + \delta \mathcal{X}(s+\theta)) \right] \\ &= \frac{a_0^2}{2} E \left[ \cos(2\pi f_0(t-s) + \delta [\mathcal{X}(t+\theta) - \mathcal{X}(s+\theta)]) \right. \\ &\quad \left. + \cos(2\pi f_0(t+s+2\theta) + \delta (\mathcal{X}(t+\theta) + \mathcal{X}(s+\theta))) \right] \rightarrow 0 \text{ for same reasons above} \\ &= \frac{a_0^2}{2} E \left[ \cos(2\pi f_0(t-s) + \delta [\mathcal{X}(t+\theta) - \mathcal{X}(s+\theta)]) \right] \end{aligned}$$

If  $|t-s| > T$ , then

$$\begin{aligned} &= \frac{a_0^2}{2} \left[ \frac{1}{4} \cos(2\pi f_0(t-s) + 2\delta) + \frac{1}{2} \cos(2\pi f_0(t-s)) + \frac{1}{4} \cos(2\pi f_0(t-s) - 2\delta) \right] \\ &= \frac{a_0^2}{2} \left[ \frac{1}{2} \cos(2\pi f_0(t-s)) \cos(2\delta) + \frac{1}{2} \cos(2\pi f_0(t-s)) \right] \\ &= \frac{a_0^2}{4} \left[ \cos(2\pi f_0(t-s)) (1 + \cos(2\delta)) \right] \end{aligned}$$

If  $|t-s| \leq T$ , then

$$= \frac{a_0^2}{2} \left( 1 - \frac{|t-s|}{T} \right) \cos(2\pi f_0(t-s)) + \frac{|t-s|}{T} \frac{a_0^2}{4} \left[ \cos(2\pi f_0(t-s)) (1 + \cos(2\delta)) \right]$$

This results in an autocorrelation function

$$R_V(t,s) = \begin{cases} \frac{q_0^2}{2} \left(1 - \frac{|t-s|}{T}\right) \cos(2\pi f_0(t-s)) + \frac{|t-s|}{T} \left(\frac{q_0^2}{4}\right) \cos(2\pi f_0(t-s)) (1 + \cos(2\delta)) & , |t-s| \leq T \\ \frac{q_0^2}{4} \cos(2\pi f_0(t-s)) (1 + \cos(2\delta)) & , |t-s| > T \end{cases}$$

Notice that  $R_X(t,s)$  and  $R_V(t,s)$  are dependent only on the time difference, i.e. they are weakly stationary. Thus, we could calculate the PSD

from the Fourier transform of the autocorrelation



5.29)  $\{\bar{X}_n\}, \{\bar{Z}_n\}$  are zero-mean, mutually independent, iid, two-sided Gaussian random processes.

$$R_{\bar{X}}(k) = \sigma_{\bar{X}}^2 \delta_k; \quad R_{\bar{Z}}(k) = \sigma_{\bar{Z}}^2 \delta_k$$

$$Y_n = Z_n + r Y_{n-1}$$

$$U_n = \bar{X}_n + Z_n$$

$$W_n = U_n + r U_{n-1}$$

$$K_U(n, k) = E[(U_n - \bar{U}_n)(U_k - \bar{U}_k)]$$

$$= E[(\bar{X}_n + Z_n - \bar{X}_n - \bar{Z}_n)(\bar{X}_k + Z_k - \bar{X}_k - \bar{Z}_k)]$$

$$= K_{\bar{X}}(n, k) + K_{\bar{Z}}(n, k) + E[(\bar{X}_n - \bar{X}_n)(Z_k - \bar{Z}_k)] + E[(\bar{X}_k - \bar{X}_k)(Z_n - \bar{Z}_n)]$$

$$= K_{\bar{X}}(n, k) + K_{\bar{Z}}(n, k) + E[\cancel{\bar{X}_n - \bar{X}_n}] E[\cancel{Z_k - \bar{Z}_k}] + E[\cancel{\bar{X}_k - \bar{X}_k}] E[\cancel{Z_n - \bar{Z}_n}]$$

$$= R_{\bar{X}}(n, k) + R_{\bar{Z}}(n, k)$$

$$= \sigma_{\bar{X}}^2 \delta(n-k) + \sigma_{\bar{Z}}^2 \delta(n-k) = \sigma_{\bar{X}}^2 + \sigma_{\bar{Z}}^2 \delta(n-k)$$

$$S_U = \sigma_{\bar{X}}^2 + \sigma_{\bar{Z}}^2$$

$$K_W(n, k) = E[(W_n - \bar{W}_n)(W_k - \bar{W}_k)]$$

$$= E[(U_n + r U_{n-1} - \bar{U}_n - r \bar{U}_{n-1})(U_k + r U_{k-1} - \bar{U}_k - r \bar{U}_{k-1})]$$

$$= E[(U_n - \bar{U}_n)(U_k - \bar{U}_k)] + r^2 E[(U_{n-1} - \bar{U}_{n-1})(U_{k-1} - \bar{U}_{k-1})]$$

$$+ E[(U_n - \bar{U}_n)(U_{k-1} - \bar{U}_{k-1})] + r^2 E[(U_{n-1} - \bar{U}_{n-1})(U_k - \bar{U}_k)]$$

$$= K_U(n, k) + r^2 K_U(n-1, k-1) + K_U(n, k-1) + r^2 K_U(n-1, k)$$

$$= (\sigma_{\bar{X}}^2 + \sigma_{\bar{Z}}^2) (\delta(n-k) + r^2 \delta(n-k) + \delta(n-k+1) + r^2 \delta(n-1-k))$$

$$= (\sigma_{\bar{X}}^2 + \sigma_{\bar{Z}}^2) (\delta(n-k) + r^2 \delta(n-k) + \delta(n-k+1) + r^2 \delta(n-k-1))$$

$$K_W(k) = (\sigma_{\bar{X}}^2 + \sigma_{\bar{Z}}^2) (\delta(k) + r^2 \delta(k) + \delta(k+1) + r^2 \delta(k-1))$$

$$S_W(f) = (\sigma_{\bar{X}}^2 + \sigma_{\bar{Z}}^2) (1 + r^2 + e^{-j2\pi f(-1)} + r^2 e^{-j2\pi f(+1)})$$

$$E[(X_n - W_n)^2] = E[(X_n - U_n - rU_{n-1})^2]$$

$$= E[(X_n - (X_n - Z_n) - r(X_{n-1} - Z_{n-1}))^2]$$

$$= E[(Z_n - rX_{n-1} + rZ_{n-1})^2]$$

$$= E[Z_n^2 - rZ_nX_{n-1} + rZ_nZ_{n-1} - rZ_nX_{n-1} + r^2X_{n-1}^2 - r^2X_{n-1}Z_{n-1} + rZ_nZ_{n-1} - r^2X_{n-1}Z_{n-1} + r^2Z_{n-1}^2]$$

$$= E[Z_n^2 + r^2X_{n-1}^2 + r^2Z_{n-1}^2]$$

$$= \sigma_z^2 + r^2\sigma_x^2 + r^2\sigma_z^2$$

$$= \underline{r^2\sigma_x^2 + (1+r^2)\sigma_z^2}$$

5.30)  $\{Z_n\}$  and  $\{W_n\}$  are two mutually independent two-sided zero-mean iid Gaussian processes with variances  $\sigma_z^2$  and  $\sigma_w^2$ .

$$\bar{X}_n = Z_n - rZ_{n-1} \quad \text{with } 0 < r < 1.$$

$$Y_n = \bar{X}_n + W_n$$

$$U_n = rU_{n-1} + Y_n$$

a)  $Z_n$  is iid.  $R_Z(k) = \sigma_z^2 \delta_k \therefore S_Z(f) = \sigma_z^2$  all  $f$

b)  $R_{\bar{X}}(k) = E[(Z_n - rZ_{n-1})(Z_{n+k} - rZ_{n+k-1})]$   
 $= R_Z(k) - rR_Z(k-1) - rR_Z(k+1) + r^2R_Z(k)$   
 $= \sigma_z^2 [(1+r^2)\delta_k - r(\delta_{k-1} + \delta_{k+1})]$

$$S_{\bar{X}}(f) = \sigma_z^2 \left( (1+r^2) - r(e^{-j2\pi f} + e^{j2\pi f}) \right)$$

$$= \sigma_z^2 (1+r^2 - 2r \cos(2\pi f))$$

c)  $R_Y(k) = E[Y_n Y_{n+k}]$   
 $= E[(\bar{X}_n + W_n)(\bar{X}_{n+k} + W_{n+k})]$   
 $= R_{\bar{X}}(k) + \cancel{E[W_n \bar{X}_{n+k}]} + \cancel{E[\bar{X}_n W_{n+k}]} + R_W(k)$   
 $= \sigma_z^2 [(1+r^2)\delta_k - r(\delta_{k-1} + \delta_{k+1})] + \sigma_w^2 \delta_k$

$$S_Y(f) = \sigma_z^2 (1+r^2 - 2r \cos(2\pi f)) + \sigma_w^2$$

d)  $E[(U_n - Z_n)^2]$

$$U_n = rU_{n-1} + \bar{X}_n + W_n$$

$$U_n = rU_{n-1} + Z_n - rZ_{n-1} + W_n$$

$$U_n - Z_n = rU_{n-1} + Z_n - rZ_{n-1} + W_n - Z_n$$

$$U_n - Z_n = rU_{n-1} - rZ_{n-1} + W_n$$

from above

$$\begin{aligned}
 U_n &= rU_{n-1} + \bar{X}_n + W_n \\
 &= r(rU_{n-2} + \bar{X}_{n-1} + W_{n-1}) + \bar{X}_n + W_n \\
 &= r(r(rU_{n-3} + \bar{X}_{n-2} + W_{n-2}) + \bar{X}_{n-1} + W_{n-1}) + \bar{X}_n + W_n \\
 &= r(r(r(rU_{n-4} + \bar{X}_{n-3} + W_{n-3}) + \bar{X}_{n-2} + W_{n-2}) + \bar{X}_{n-1} + W_{n-1}) + \bar{X}_n + W_n \\
 &= r^4 U_{n-4} + r^3 \bar{X}_{n-3} + r^3 W_{n-3} + r^2 \bar{X}_{n-2} + r^2 W_{n-2} + r \bar{X}_{n-1} + r W_{n-1} + \bar{X}_n + W_n \\
 &= r^k U_{n-k} + \sum_{i=0}^k r^i \bar{X}_{n-i} + \sum_{i=0}^k r^i W_{n-i} \\
 &= r^k U_{n-k} + \sum_{i=0}^k r^i (Z_{n-i} - rZ_{n-i-1}) + \sum_{i=0}^k r^i W_{n-i}
 \end{aligned}$$

This is a two-sided process so  $k \rightarrow \infty$

$$\begin{aligned}
 U_n &= \sum_{i=0}^{\infty} r^i (Z_{n-i} - rZ_{n-i-1}) + \sum_{i=0}^{\infty} r^i W_{n-i} \\
 &= \sum_{i=0}^{\infty} r^i Z_{n-i} - r \sum_{i=0}^{\infty} r^i Z_{n-i-1} + \sum_{i=0}^{\infty} r^i W_{n-i}
 \end{aligned}$$

$$U_n - Z_n = \sum_{i=0}^{\infty} r^i Z_{n-i} - r \sum_{i=0}^{\infty} r^i Z_{n-i-1} + \sum_{i=0}^{\infty} r^i W_{n-i} - Z_n$$

$$\begin{aligned}
 E[(U_n - Z_n)^2] &= E\left[\left(\sum_{i=0}^{\infty} r^i Z_{n-i} - r \sum_{i=0}^{\infty} r^i Z_{n-i-1} + \sum_{i=0}^{\infty} r^i W_{n-i} - Z_n\right)^2\right] \\
 &= E\left[\left(\sum_{i=0}^{\infty} r^i Z_{n-i} - \sum_{i=0}^{\infty} r^i Z_{n-i}\right) + \sum_{i=0}^{\infty} r^i W_{n-i}\right]^2
 \end{aligned}$$

Because  $W$  are mutually independent

$$E\left[\left(\sum_{i=0}^{\infty} r^i W_{n-i}\right)^2\right] = \sum_{i=0}^{\infty} r^{2i} E[W_n^2] = \boxed{\frac{\sigma_W^2}{1-r^2}}$$

