

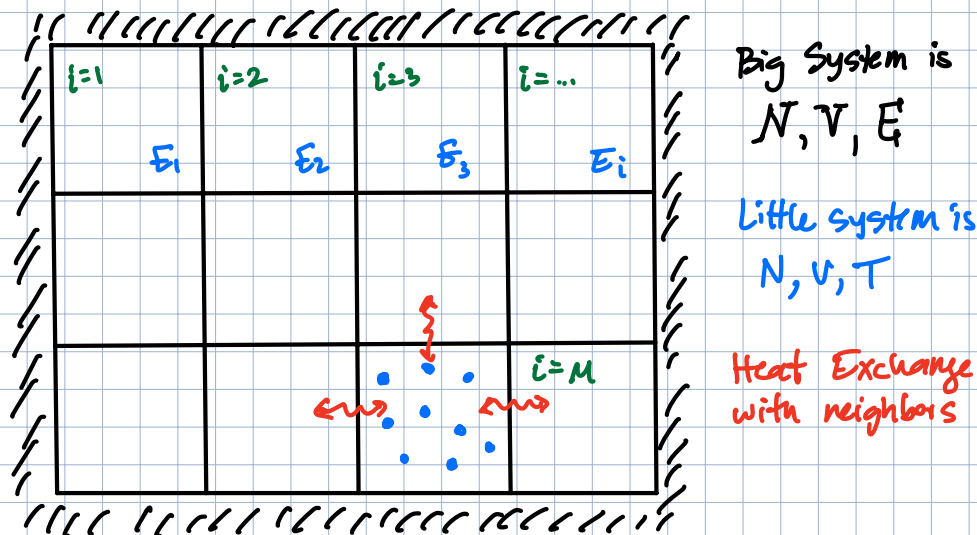
Lecture 17 - The Canonical Ensemble

A. Introduction to the Canonical Ensemble

We saw that we could define an ensemble of systems that had a fixed number of particles N , volume V , and energy E . However, energy and temperature are thermodynamic conjugates. E is not privileged over T .

Also, it is more common for us to experimentally encounter a system that is at constant T . So, we would like to construct an ensemble with constant N , V , and T . This type of ensemble is called a canonical ensemble.

We are going to construct a canonical ensemble as follows. We want to use a microcanonical ensemble so we don't need any new postulates or assumptions.



- There are M "little systems" making up the ensemble.

$$N = M N, \quad V = M V$$

- The other systems act as a heat bath for the NVT system, keeping it at a constant temperature.

- The energy of the NVT system fluctuates and is a random variable E_i .
- In the M members of the ensemble, there will be m_j that have energy E_j .

$$m_j = \sum_{i=1}^M \delta(E_i, E_j)$$

$\uparrow$
 +1 if $E_i = E_j$

m_j : occupation number

J : number of energy levels

$$\delta(E_i, E_j) = \begin{cases} 1 & \text{if } E_i = E_j \\ 0 & \text{otherwise} \end{cases}$$

- The total energy and total number of ensembles are constrained

$$\sum_j m_j E_j = E \quad \underline{m} = \{m_j\}, \text{ vector of all occupation \#s}$$

$$\sum_j m_j = M$$

There are many different ways the energy of the microcanonical ensemble E can be distributed in the M NVT systems

Example: $M=6$ $E=8$ $J=4$ $E_j = \{0, 1, 2, 3\}$

Distribution 1

i	1	2	3	4	5	6
E_i	2	1	1	2	1	1

$$\underline{m} = [0, 4, 2, 0]$$

$$\sum_j m_j = 0 + 4 + 2 + 0 = 6 = M \checkmark$$

$$\sum_j m_j E_j = 0 \cdot 0 + 4 \cdot 1 + 2 \cdot 2 + 0 \cdot 3 = 4 + 4 = 8 = E \checkmark$$

Distribution 2

i	1	2	3	4	5	6
E_i	0	3	1	2	1	1

$$\underline{m} = [1, 3, 1, 1]$$

$$\sum_j m_j = 1 + 3 + 1 + 1 = 6 = M \checkmark$$

$$\sum_j m_j E_j = 1 \cdot 0 + 3 \cdot 1 + 1 \cdot 2 + 1 \cdot 3 = 3 + 2 + 3 = 8 = E \checkmark$$

B. Probability Density for the Canonical Ensemble

For the microcanonical ensemble, we

1. Postulated a probability density: uniform density of configurations with equal energy.
2. Used the probability density to find the density of states $\Omega(E)$.
3. Used the Boltzmann formula to find the entropy: $S = k_B \ln \Omega$

We want to follow an analogous procedure for the canonical ensemble. Let's first find the probability density of finding one of the members of our NVT ensemble at energy E_j

$$P_j = P_{NVT}(E_j) = ?$$

If we could look at the statistics of our ensemble, we would see that

$$P_j = \frac{\langle m_j \rangle}{M} \quad \begin{array}{l} \leftarrow \text{average \# of times we observe} \\ \text{an NVT ensemble w/ energy } E_j \text{ in NVE} \quad (*) \\ \leftarrow \text{total \# of NVT ensembles} \end{array}$$

In the microcanonical ensemble, we said that any configuration with energy E was equally probable (principle of equal a priori probability). Here, the big system is an NVE ensemble. That means that any vector $\underline{m}$ where $\sum_j m_j E_j = E$ is equally probable. So, we just to count the number of ways we can arrange the energy in the M bins. This will tell us the probability distribution.

$W(\underline{m})$, # of ways of arranging the states for a given $\underline{m}$

$$W(\underline{m}) = \frac{M!}{m_1! m_2! \dots m_j!} = \frac{M!}{\prod_{j=1}^J m_j!} \quad \text{Multinomial coefficient}$$

Using $W(\underline{m})$ we can write a probability mass function for a configuration,

$$P_{NVE}(\underline{m}) = \frac{W(\underline{m})}{\sum_{\underline{m}} W(\underline{m})}$$

PMF of the configuration $\underline{m}$ in the NVE ensemble (multidimensional of all $\underline{m}$)

Using this probability mass function we can find the average

$$P_{NVE}(E_j) = \frac{\langle m_j \rangle}{M} = \frac{1}{M} \sum_{\underline{m}} P_{NVE}(\underline{m}) m_j(\underline{m}) = \frac{1}{M} \frac{\sum_{\underline{m}} W(\underline{m}) m_j(\underline{m})}{\sum_{\underline{m}} W(\underline{m})}$$

Rewriting (*) using $W(\underline{m})$

Example: $M=6$ $E_T=8$ $J=4$ $E_j = \{0, 1, 2, 3\}$

Distribution 1: $\underline{m} = [0, 4, 2, 0]$

$$W(\underline{m}) = \frac{6!}{0! 4! 2! 0!} = \frac{3 \cdot \cancel{6} \cdot 5 \cdot \cancel{4} \cdot \cancel{3} \cdot \cancel{2} \cdot 1}{\cancel{4} \cdot \cancel{3} \cdot \cancel{2} \cdot \cancel{1}} = 15$$

of ways to arrange the $4 \times E=1$ and $2 \times E=2$ into 6 bins.

Distribution 2: $\underline{m} = [1, 3, 1, 1]$

$$W(\underline{m}) = \frac{6!}{1! 3! 1! 1!} = \frac{6 \cdot 5 \cdot 4 \cdot 3 \cdot 2 \cdot 1}{3 \cdot 2 \cdot 1} = 6 \cdot 5 \cdot 4 = 120$$

Many more ways to come up with this distribution of energies

note: $\sum_{\underline{m}} W(\underline{m}) =$ Add up all possible configurations (Hard to write as simple sum b/c of constraints)

$m_j(\underline{m})$: # of ensembles with energy E_j when configuration is $\underline{m}$.

($m_0 = 0$ in Dist 1 and $m_4 = 1$ in Dist 2.)

Show the Python code for M ensembles with $J=4$ energy levels and fixed energy $E_T=8$.

Question: Is W the same as the density of states Ω ? No, because it is related to the distribution of the microstates. It is related to S in NVT, which has a T -dependence.

Ok, this is all great, but what do I do when $M \rightarrow \infty$ (large ensemble) and J is big (large number of energies). Computing the multinomial coefficient will be very hard. Answer: when the $m_j \rightarrow \infty$, the multinomial coefficient becomes a very sharply peaked normal distribution around $m_j = m_j^*$. So sharp that it becomes a Dirac delta function $\delta(m_j - m_j^*)$ and only the most probable configuration matters. ** This maximizes the entropy of the NVE system.*

Because the distribution is so sharp, the average is the same as the most likely state m_j^* . Therefore, our probability becomes

$$\langle m_j \rangle = m_j^* \Rightarrow P_{NVT}(E_j) = \frac{m_j^*}{M}$$

Example: Show that $\langle m_j \rangle = m_j^*$ when $M \rightarrow \infty$ for $J=2$ using Python. Proof is in the Appendix.

C. The Boltzmann Distribution

We want to find the most probable configuration for the multi-dimensional $W(m)$ subject to the constraints

$$\sum_j m_j = M \quad \leftarrow \text{total \# of ensembles}$$

$$\sum_j m_j E_j = E_t \quad \leftarrow \text{total energy}$$

We will use the method of Lagrange multipliers for the constraints. Let

$$f(m) = \ln W(m) - \alpha \left(\sum_j m_j - M \right) - \beta \left(\sum_j m_j E_j - E_t \right)$$

Find the extremum:

$$\left. \frac{df}{dm_j} \right|_{m=m^*} = 0$$

maximizing $W(m)$ is the same as maximizing $P(m)$ because $P = W(m)/\text{const.}$ Easier to maximize the log, too.

$$\frac{df}{dm_j} = \underbrace{\frac{d}{dm_j}(\ln W)}_1 - \alpha \underbrace{\frac{d}{dm_j}(\sum_j m_j - M)}_2 - \beta \underbrace{\frac{d}{dm_j}(\sum_j m_j E_j - E)}_3$$

$$\begin{aligned} 1: \frac{d}{dm_j}(\ln W) &= \frac{d}{dm_j} \ln \left(\frac{M!}{m_1! m_2! \dots m_j! \dots m_J!} \right) \quad \leftarrow \text{Don't need to sub for } m_j \text{ because of constraints} \\ &= \frac{d}{dm_j}(\ln M!) - \frac{d}{dm_j}(\ln m_1!) - \frac{d}{dm_j}(\ln m_2!) - \dots - \frac{d}{dm_j}(\ln m_j!) \\ &\quad - \dots - \frac{d}{dm_j}(\ln m_J!) \\ &= -\frac{d}{dm_j}(\ln m_j!) = -\frac{d}{dm_j}(m_j \ln m_j - m_j) = -(\ln m_j + \cancel{\frac{m_j}{m_j}} - 1) \\ &= -\ln m_j \end{aligned}$$

$$2: \frac{d}{dm_j}(\sum_j m_j - M) = \frac{d}{dm_j}(m_1 + m_2 + \dots + m_j + \dots + m_{J-1} + m_J - M) = 1$$

$$\begin{aligned} 3: \frac{d}{dm_j}(\sum_j m_j E_j - E) &= \frac{d}{dm_j}(m_1 E_1 + m_2 E_2 + \dots + m_j E_j + \dots + m_{J-1} E_{J-1} + m_J E_J - E) \\ &= E_j \end{aligned}$$

Put it together:

$$\left. \frac{df}{dm_j} \right|_{m_j = m_j^*} = -\ln m_j^* - \alpha - \beta E_j = 0$$

$$m_j^* = e^{-\alpha} e^{-\beta E_j}$$

ok, finally, let's get the distribution

$$P(E_j) = \frac{m_j^*}{M} = \frac{e^{-\alpha} e^{-\beta E_j}}{M}$$

Summing over all the systems allows us to find α ,

$$\sum_{i=1}^M P(E_i) = 1$$

$$\sum_{i=1}^M \frac{e^{-\alpha} e^{-\beta E_i}}{M} = 1 \Rightarrow \frac{e^{-\alpha}}{M} \sum_{i=1}^M e^{-\beta E_i} = 1 \Rightarrow e^{-\alpha} = M \sum_{i=1}^M e^{-\beta E_i}$$

Putting it all together gives us the probability density!

$$P_{NVT}(E_i) = \frac{e^{-\beta E_i}}{\sum_{i=1}^M e^{-\beta E_i}}$$

We have yet to identify β .
We will see that $\beta = 1/k_B T$.

← Boltzmann Distribution!

Summary of derivation

- We create an ensemble of NVT distributions that gives a total NVE system.
- Each NVT system has a random energy E_j .
- Any distribution of the energies, E_j that sums to E is equally likely to occur, but some of them occur many ways.
- For a large ensemble, only the most likely way to distribute the energies (the one with the most ways) will occur.
- We can solve this by maximizing the system, keeping the # of ensembles and energy constant.
- This gives the Boltzmann distribution.

D. Appendix: Proof that $\mathcal{P}_{NVE}$ becomes $\delta(m-m^*)$ at $M \rightarrow \infty$

Assume Binomial Coefficient ($J=2$) to make tractable.

$$W(m) = \frac{M!}{m_1! m_2!} \quad m_1 = m \quad m_2 = M-m$$

$$= \frac{M!}{m! (M-m)!} = \binom{M}{m}$$



(i.) Determining the maximum of the distribution

The maximum of $\ln W$ is the same as the maximum of W .

$$\begin{aligned} \ln W &= \ln M! - \ln m! - \ln (M-m)! \\ &= M \ln M - M - (m \ln m - m) - [(M-m) \ln (M-m) - (M-m)] \\ &= M \ln M - \cancel{M} - m \ln m + \cancel{m} - M \ln (M-m) \\ &\quad + m \ln (M-m) + \cancel{M} - \cancel{m} \\ &= M \ln \left(\frac{M}{M-m} \right) - m \ln \left(\frac{m}{M-m} \right) \end{aligned}$$

Now to find the maximum, we look for the extremum

$$\begin{aligned} \frac{d}{dm} (\ln W) \Big|_{m^*} &= 0 \\ \frac{d}{dm} \left[\cancel{M} \ln M - M \ln (M-m) - m \ln m + m \ln (M-m) \right] &= 0 \\ -\frac{M}{M-m} \cdot (-1) - \ln m - \frac{m}{m} + \ln (M-m) \cdot (-1) + \frac{m}{M-m} (-1) &= 0 \\ \underbrace{-\frac{M}{M-m} \cdot (-1)}_{+1} - \ln m - \underbrace{\frac{m}{m}}_1 + \underbrace{\ln (M-m) \cdot (-1)}_{\text{combine}} + \underbrace{\frac{m}{M-m} (-1)}_{\text{combine}} &= 0 \\ \underbrace{\frac{(M-m)}{M-m}}_{+1} + \ln \left(\frac{M-m}{m} \right) - 1 &= 0 \Rightarrow \frac{M-m}{m} = e^0 = 1 \\ M-m = m \Rightarrow 2m = M \Rightarrow \boxed{m^* = \frac{M}{2}} \end{aligned}$$

(ii.) Determine the distribution as $M \rightarrow \infty$

The distribution (PMF) is

$$P(m) = \frac{\binom{M}{m}}{\sum_m \binom{M}{m}} = \frac{1}{2^M} \binom{M}{m}$$

To find the distribution at large M , we will look at the cumulant generating function. Find the probability generating function first

$$G(z) = E[z^m]$$

$$G(z) = \sum_{m=0}^M \frac{1}{2^M} \binom{M}{m} z^m = \frac{1}{2^M} (1+z)^M = \left(\frac{1+z}{2}\right)^M$$

binomial series

The moment generating function is

$$M(s) = E(e^{sm}) = G(e^s) = \left(\frac{1+e^s}{2}\right)^M$$

The cumulant generating function is

$$K(s) = \ln M(s) = \ln \left[\left(\frac{1+e^s}{2}\right)^M \right] = M \ln(1+e^s) - M \ln 2$$

Now find cumulants

$$k_1 = \mu = \left. \frac{dK}{ds} \right|_{s=0} = \left. \frac{M}{1+e^s} e^s \right|_{s=0} = \frac{M(1)}{1+1} = M/2$$

the maximum!
 $m^* = M/2$

$$k_2 = \sigma^2 = \left. \frac{d^2 K}{ds^2} \right|_{s=0} = \left[M e^s \cdot \left(\frac{1}{1+e^s}\right) + \frac{-M e^s}{(1+e^s)^2} e^s \right]_{s=0} = \left[\frac{M}{2} - \frac{M}{2^2} \right] = \frac{M}{4}$$

(used Mathematica for $k_3, \dots$)

$$k_3 = 0 \quad k_4 = -M/8 \quad k_5 = 0 \quad k_6 = M/4$$

Series solution

$$K(s) = \frac{M}{2} s + \frac{M}{4} \frac{s^2}{2!} - \frac{M}{8} \frac{s^4}{4!} + \frac{M}{4} \frac{s^6}{6!} + \dots$$

Compare to normal (Gaussian):

$$K_{\text{Gauss}} = \mu S + \frac{\sigma^2}{2} S^2$$

Do the higher order terms matter as $M \rightarrow \infty$? (No) S is of the order of $1/\sqrt{M}$ (e^{SM}). let's rescale.

$$\text{let } \tilde{S} = MS \rightarrow S = \tilde{S}/M$$

$$K(\tilde{S}) = \frac{M}{2} \frac{\tilde{S}}{M} + \frac{M}{4} \frac{\tilde{S}^2}{M^2} - \frac{M}{8} \cdot \frac{\tilde{S}^4}{24 M^4} + \frac{M}{4} \frac{\tilde{S}^6}{720 M^6}$$

$$K(\tilde{S}) = \frac{1}{2} \tilde{S} + \frac{1}{8M} \tilde{S}^2 - \frac{1}{192 M^3} \tilde{S}^4 + \frac{1}{2880 M^5} \tilde{S}^6$$

The width
will also go
to 0 like $M^{-1/2}$.

These terms will drop out
very quickly as $M \rightarrow \infty$

Conclusion: By CLT $K(S) \rightarrow$ Gaussian with a maximum at $M/2$ and width $\sqrt{M/4}$. It will get sharper and sharper as $M \rightarrow \infty$.

(ii). More accurate Stirling's Approximation

For plotting the Stirling's approximation, $O(\ln M)$ corrections make a difference. let's do this one out real quick

$$W = \frac{M!}{m!(M-m)!} \quad \ln n! \sim n \ln n - n + \frac{1}{2} \ln(2\pi n)$$

$$\ln W = \ln M! - \ln m! - \ln(M-m)!$$

$$= M \ln M - M + \frac{1}{2} \ln(2\pi M) - \left[m \ln m - m + \frac{1}{2} \ln(2\pi m) \right] \\ - \left[(M-m) \ln(M-m) - (M-m) + \frac{1}{2} \ln(2\pi(M-m)) \right]$$

$$= \underbrace{M \ln M}_{\text{orange}} - \cancel{M} + \frac{1}{2} \ln(2\pi M) - \underbrace{m \ln m}_{\text{green}} + \cancel{m} - \frac{1}{2} \ln(2\pi m) \\ - \underbrace{M \ln(M-m)}_{\text{orange}} + \cancel{M} - \frac{1}{2} \ln(2\pi(M-m)) + \underbrace{m \ln(M-m)}_{\text{green}} - \cancel{m}$$

$$= M \ln \left(\frac{M}{M-m} \right) - m \ln \left(\frac{m}{M-m} \right) + \frac{1}{2} \ln \left[\frac{2\pi M}{(2\pi m)(2\pi(M-m))} \right]$$

$$= M \ln \left(\frac{M}{M-m} \right) - m \ln \left(\frac{m}{M-m} \right) + \frac{1}{2} \ln \left(\frac{1}{2\pi} \frac{M}{M-m} \right)$$

This extra term helps the approximation at moderate m .

The normalization constant for the exact binomial coefficient is 2^M .

In other words

$$\sum_{m=0}^M \binom{M}{m} = 2^M$$